#### **Greek Letters**

 $\alpha$  = complementary apex angle

 $\dot{\gamma}$  = shear rate  $\omega$  = angular velocity

 $\eta = \text{viscosity}$  = apparent vi

 $\eta_a = \text{apparent viscosity}$   $\eta_o = \text{zero-shear viscosity}$ 

 $\eta_{oM}$ ,  $\eta_{oL} =$  zero-shear viscosity of the more and less vis-

cous components, respectively

λ = characteristic time

 $\sigma_1$  = primary normal stress coefficient

 $\theta_M$  = mixing time

 $\theta_{B}$  = blending time = density

Manuscript received April 25, 1975; revision received August 8, and accepted August 11, 1975.

## LITERATURE CITED

Coyle, C. K., H. E. Hirschland, B. J. Michel and J. Y. Oldshue, AIChE J., 16, 906 (1970).

Ford, D. E., R. A. Mashelkar, and J. Ulbrecht, "Mixing Times in Newtonian and Non-Newtonian Fluids," Proc. Tech. Int., 17, 803-806 (1972).

Nagata, S., M. Yanagimoto, and T. Yokoyama, Mem. Fac. Eng. Kyoto Univ., 18, 444 (1956).

# Probability Models in Reaction Path Synthesis

## L. T. FAN and R.H. WANG

Department of Chemical Engineering Kansas State University, Manhattan, Kansas 66506

A significant contribution has been made in the field of reaction path synthesis by Powers and his associates (1975). In employing their approach, we need to know the probability of failure in each step of the reaction path and the number of trials required to obtain the first success at each step. This note introduces a general probability model for the reaction path synthesis which can be used to predict the number of trials required to obtain the first success for sequential trials and at least one success for simultaneous tests.

#### **FORMULATION**

As shown in Figure 1, n sequential trials are conducted in the horizontal direction in synthesizing a reaction path. At the i<sup>th</sup> trial the reagent is divided into  $l_i$  portions,  $M_{i1}$ ,  $M_{i2}$ , ...,  $M_{il_i}$ , distributed in the vertical direction, which can be used either simultaneously or sequentially. Several specific cases can be considered.

Case 1. Sequential both in the horizontal and vertical directions. The probability for the first success to occur at the  $M_{ij}^{th}$  test (the  $j^{th}$  test at the  $i^{th}$  trial) can be expressed as (Parzen, 1960)

Pr {the first success to occur at the  $M_{ii}$ <sup>th</sup> test}

$$= (1 - P_s)^{l_1} (1 - P_s)^{l_2} \dots (1 - P_s)^{l_{i-1}} (1 - P_s)^{j-1} P_s$$

$$= (1 - P_s)^{\sum_{r=1}^{i-1} (l_r) + j - 1} P_s \tag{1}$$

where  $P_s$  is the probability of success in a single test. Equation (1) is the geometrical probability density function. The expected number of tests to obtain the first success, according to this probability density function, is (Parzen, 1962; Appendix 1).

$$E\left(\sum_{r=1}^{i-1} (l_r) + j\right) = \frac{1}{P_s} \tag{2}$$

If we have

$$l_1=l_2=\ldots=l_i=1$$

and consequently i = 1, we have from Equation (1)

Pr {the first success to occur at the ith trial}

$$= (1 - P_s)^{i-1}P_s$$

which is the geometrical probability density function in a simple form (Parzen, 1960). Of course, the expected num-

Fig. 1. A probability model for the reaction path synthesis.

ber of trials to obtain the first success can still be calculated by Equation (2).

Case 2. Sequential in the horizontal direction and simultaneous in the vertical direction. For this case, we may have more than one success at any given trial because  $l_i$  tests are carried out simultaneously at the i<sup>th</sup> trial. The probability for at least one success to occur at the i<sup>th</sup> trial can be expressed as (Feller, 1950; Parzen, 1962)

Pr {at least one success in  $l_i$  simultaneous tests to occur at the ith trial}

$$= (1 - P_s)^{\frac{l-1}{\sum l_r}} \{1 - (1 - P_s)^{l_i}\}$$

$$i \ge 1$$
(3)

where

t 
$$Q_{\mathfrak{i}} \equiv (1-P_s)^{l_{\mathfrak{i}}}$$
 where  $Q_{\mathfrak{0}} \equiv (1-P_s)^{l_{\mathfrak{0}}} \equiv 1$ 

 $Q_i$  is the probability that the i<sup>th</sup> trial is a failure. Thus

$$R_i = 1 - Q_i$$

 $l_0 \equiv 0$ 

is the probability that the i<sup>th</sup> trial is a success. Then Equation (3) can be rewritten as

Pr {at least one success to occur at the i<sup>th</sup> trial}

$$= \begin{pmatrix} i-1 \\ \pi \\ r=0 \end{pmatrix} \cdot R_i \quad i \ge 1 \quad Q_0 \equiv 1 \tag{4}$$

Thus the expected number of trials to obtain at least one success at the  $i^{\rm th}$  trial is (Appendix II)\*

$$E(i)' = \sum_{i=1}^{\infty} i {i \choose \frac{\pi}{r=0}} Q_r \cdot R_i$$

$$= 1 + Q_1 + Q_1 Q_2 + Q_1 Q_2 Q_3 + \dots + Q_1 Q_2 \dots Q_i + \dots$$
 (5)

This case can be resolved into the following subcases: Case 2.1. If

$$l_1 = l_2 = \ldots = l_i = l$$

we have, according to Equation (3)

Pr {at least one success to occur at the ith trial}

$$= (1 - P_s)^{l(i-1)} - (1 - P_s)^{li}$$
 (6)

and thus (Appendix III)\*

$$E(i) = \frac{1}{1 - (1 - P_s)^t} \tag{7}$$

Case 2.2. If i = 1, according to Equation (3), we have

 $p_1 = Pr$  (at least one success in  $l_1$  simultaneous tests to occur at the first trial}

$$=1-(1-P_s)^{l_1} (8)$$

or

$$1 - p_1 = (1 - P_s)^{l_1} (9)$$

Equation (8) can also be derived from the binomial probability density function which is (Parzen, 1962)

Pr {exactly k successes to occur in  $l_1$  simultaneous tests at the first trial}

$$= \binom{l_1}{k} P_s^k (1 - P_s)^{l_1 - k}$$

Then

 $p_1 = Pr$  {at least one success in  $l_1$  simultaneous tests to occur at the first trial}

= 1 - Pr (exactly zero successes to occur in  $l_1$  simultaneously tests at the first trial}

$$=1-(1-P_s)^{l_1}$$

which is Equation (8).

$$p_0 = 1 - p_1$$

where  $p_0$  is the probability of exactly zero successes to occur in  $\bar{l_1}$  simultaneous tests. Since the probability of failure

$$P_F = 1 - P_s$$

Equation (9) can be rewritten as

$$p_0 = P_F^{l_1} \tag{10}$$

Then  $l_1$  is given by solving Equation (10) as

$$l_1 = \frac{\ln p_0}{\ln P_F} \tag{11}$$

This is the number of tests required to achieve at least one success in simultaneous testing, provided that the probability of zero successes is known to be  $p_0$  (Hadley, 1967). However, it should not be stated as the number of trials to the first success with probability of success of  $(1 - p_0)$  (Powers et al., 1975).

## ESTIMATION OF PROBABILITY OF ZERO SUCCESSES IN BINOMIAL TRIALS

The probability of exactly zero successes in n binomial trials is an important parameter in the synthesis of fault tolerant reaction paths (Powers et al., 1975). In this section, we consider the problem of obtaining a best point estimate of the probability of exactly zero successes in nbinomial trials (simultaneous tests),  $p_0$ . An approach which is similar to that of Rutemiller (1967) is employed here.

If  $P_s$  is the probability of success in any test, the probability of exactly zero successes in n simultaneous tests is given by

$$p_0 = (1 - P_s)^n (12)$$

Suppose that r successes have been observed in m simultaneous tests. Then the maximum likelihood estimator (mle) for  $p_0$  is

$$\hat{p}_0 = \left(1 - \frac{r}{m}\right)^n \tag{13}$$

for all values of n and m (Rutemiller, 1967; Appendix

An alternative estimator of  $p_0$ , namely, the minimum variance unbiased estimator (mvue), may be obtained directly from a theorem of Patil (1963) for the generized power series distribution as shown below (Appendix V):

$$\tilde{p}_0 = \frac{\binom{m-n}{r}}{\binom{m}{r}} \quad \text{if} \quad r \le m-n \tag{14}$$

$$p_0 = 0 if r > m - n$$

The mvue,  $\tilde{p}_0$ , does not exist for n > m. Rutmiller (1967) concluded that the mle,  $\stackrel{\wedge}{p}_0$  is generally preferable to the mvue,  $p_0$ , to estimate the probability of zero failures in binomial trials when this probability is anticipated to be greater than 0.50. This should also be true in estimating the probability of zero successes in binomial trials because Equations (13) and (14) are the same as those for estimating, respectively, the mle,  $\stackrel{\wedge}{p}_0$ , and the mvue,  $p_0$ , of the probability of zero failures.

## **ACKNOWLEDGMENT**

The authors gratefully acknowledge the assistance of Professor M. Siotani, Department of Statistics, Kansas State University, and Professors K. Osaki and H. Sayama, Okayama University, Japan.

## LITERATURE CITED

Feller, W., An Introduction to Probability Theory and Its Ap-

plication, p. 224, J. Wiley, New York (1950). Freund, J. E., Mathematical Statistics, p. 224, Prentice-Hall, Englewood Cliffs, N. J. (1962).

Hadley, G., Introduction to Probability and Statistical Decision Theory, p. 244-245, Holden-Day, San Francisco, Calif.

Parzen, E., Modern Probability Theory and Its Application, pp. 179, 218, Wiley, New York (1960).

Parzen, E., Stochastic Processes, pp. 13-15, Holden-Day, San

Francisco, Calif. (1962).
Patil, G. P., "Minimum Variance Unbiased Estimation and Certain Problems of Additive Number Theory," Annals of Math Stat., 34, 1050-1056 (1963).

Powers, G. J., D. L. Cummings, and R. N. S. Rathore, "Synthesis of Fault Tolerant Reaction Paths," AIChE J., 21, 90 (1975).

Rutemiller, H. C., "Estimation of the Probability of Zero Failures in m Binomial Trials," Amer. Stat. Asso. J., 272 (1967). Manuscript received March 7, 1975, and accepted September 3, 1975.

<sup>&</sup>lt;sup>a</sup> See document No. 02691, National Auxiliary Publications (NAPS).